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Leaf nitrogen from first principles: field evidence for adaptive variation with climate

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Abstract. Nitrogen content per unit leaf area (Narea) is a key variable in plant functional ecology and biogeochemistry. Narea comprises a structural component, which scales with leaf mass per area (LMA), and a metabolic component, which scales with Rubisco capacity. The co-ordination hypothesis, as implemented in LPJ and related global vegetation models, predicts that Rubisco capacity should be directly proportional to irradiance but should decrease with increases in ci : ca and temperature because the amount of Rubisco required to achieve a given assimilation rate declines with increases in both. We tested these predictions using LMA, leaf δ13C, and leaf N measurements on complete species assemblages sampled at sites on a north–south transect from tropical to temperate Australia. Partial effects of mean canopy irradiance, mean annual temperature, and ci : ca (from δ13C) on Narea were all significant and their directions and magnitudes were in line with predictions. Over 80 % of the variance in community-mean (ln) Narea was accounted for by these predictors plus LMA. Moreover, Narea could be decomposed into two components, one proportional to LMA (slightly steeper in N-fixers), and the other to Rubisco capacity as predicted by the co-ordination hypothesis. Trait gradient analysis revealed ci : ca to be perfectly plastic, while species turnover contributed about half the variation in LMA and Narea.

Interest has surged in methods to predict continuous leaf trait variation from environmental factors, in order to improve ecosystem models. Coupled carbon–nitrogen models require a method to predict Narea that is more realistic than the widespread assumptions that Narea is proportional to photosynthetic capacity, and/or that Narea (and photosynthetic capacity) are determined by N supply from the soil. Our results indicate that Narea has a useful degree of predictability, from a combination of LMA and ci : ca – themselves in part environmentally determined – with Rubisco activity, as predicted from local growing conditions. This finding is consistent with a “plant-centred” approach to modelling, emphasizing the adaptive regulation of traits. Models that account for biodiversity will also need to partition community-level trait variation into components due to phenotypic plasticity.
and/or genotypic differentiation within species vs. progressive species replacement, along environmental gradients. Our analysis suggests that variation in $N_{\text{area}}$ is about evenly split between these two modes.

1 Introduction

Nitrogen (N) is an essential nutrient for primary production and plant growth, and nitrogen content per unit leaf area ($N_{\text{area}}$) is a key variable in plant functional ecology and biogeochemistry. A strong correlation between leaf N and photosynthetic capacity has been observed, and is to be expected because typically almost half of the N in leaves is invested in the photosynthetic apparatus (Field and Mooney, 1986; Evans and Seemann, 1989; Evans, 1989). This component of $N_{\text{area}}$ is approximately proportional to the maximum rate of carboxylation ($V_{\text{cmax}}$) at standard temperature, also expressed per unit area (Wohlfahrt et al., 1999; Takashima et al., 2004; Kattge et al., 2009). Cell walls account for a further significant fraction of leaf N (Lampert and Northcote, 1960; Niinemets and Tenhunen, 1997; Onoda et al., 2004). Leaf mass per area (LMA) is positively correlated with cell-wall N (Onoda et al., 2004) and is used as an index of plant investment in cell-wall biomass (Reich et al., 1991; Wright and Cannon, 2001). Thus, $N_{\text{area}}$ can usefully be considered as the sum of a “metabolic” component related to $V_{\text{cmax}}$ and a “structural” component proportional to LMA. Leaves with high $V_{\text{cmax}}$ usually have high LMA, and so these two quantities can be at least partially correlated, as seen clearly (for example) in parallel vertical gradients of $V_{\text{cmax}}$ and LMA within canopies of one species (e.g. Niinemets and Tenhunen, 1997). Across different species and environments, however, there is scope for considerable independent variation in $V_{\text{cmax}}$ and LMA, implying the need to consider them separately.

Dynamic global vegetation models (DGVMs) are being extended to include interactive carbon (C) and N cycles (Thornton et al., 2007; Xu-Ri and Prentice, 2008; Zaehle and Friend, 2010), but there remain many open questions about the implementation of C–N coupling (Prentice and Cowling, 2013), including the control of leaf N content, which is treated quite differently by different models. For example, one common modelling approach predicts photosynthetic capacity from $N_{\text{area}}$, and $N_{\text{area}}$ in turn from soil inorganic N supply (e.g. Luo et al., 2004). This implies an assumption that the soil environment, and soil microbial activity in particular, are the primary controls of $N_{\text{area}}$ and photosynthetic capacity at the leaf level. An alternative assumption is that photosynthetic capacity is optimized as a function of irradiance, leaf-inner CO$_2$ concentration ($c_i$), and temperature (Haxeltine and Prentice, 1996; Dewar, 1996) — implicit in the widely used LPJ DGVM (Sitch et al., 2003) and other models derived from it, including LPJ-GUESS (Smith et al., 2001) and LPX (Prentice et al., 2011a; Stocker et al., 2013). This “plant-centred” approach embodies the idea that plant allocation processes (and thus, not soil microbial processes) determine leaf-level traits. Limited N supply, by this reasoning, should lead to the production of fewer leaves, rather than leaves with suboptimal capacity. More specifically, it is derived from a long-standing concept, the “co-ordination hypothesis”, which states that the Rubisco- and electron transport-limited rates of photosynthesis tend to be co-limiting under average daytime conditions (Chen et al., 1993; Haxeltine and Prentice, 1996; Maire et al., 2012). Co-limitation is optimal — even though mechanistically, it may be an inevitable outcome of leaf metabolism (Chen et al., 1993) — in the sense that it provides the right balance of investments in the biochemical machineries for carboxylation and electron transport. It implies that enzyme activities adjust, over relatively long periods (weeks or longer), so that co-limitation holds. An important consequence is that the predicted responses of photosynthetic traits and rates to environmental variables observed in the field (whether temporally, comparing different seasons, or spatially, comparing different environments) are substantially different from those seen in short-term laboratory experiments. Specifically, $V_{\text{cmax}}$ (and thus the metabolic component of $N_{\text{area}}$) is predicted to be directly proportional to irradiance, to decrease with increasing $c_i : c_a$, and to decrease with increasing temperature. These predictions are supported in general terms by an observed positive relationship between $N_{\text{area}}$ and irradiance (Field, 1983; Wright et al., 2005), a negative relationship between $N_{\text{area}}$ and $c_i : c_a$ (Wright et al., 2003; Prentice et al., 2011b, 2014), and (in woody evergreens at least) a negative relationship between $N_{\text{area}}$ and temperature (845 species: data from Wright et al., 2004). But there has been no systematic attempt to quantitatively assess the relationship of leaf N with environmental and structural predictors across environmental gradients. Such empirical work is needed to assess and underpin methods of C–N cycle coupling in DGVMs.

Here we set out to test the predictability of $N_{\text{area}}$ using measurements carried out on dried plant material collected by the Terrestrial Ecosystem Research Network (TERN) AusPlots and Australian Transect Network facilities, at 27 sites on a north–south transect across the Australian continent. The transect extended from the wet–dry (monsoonal) tropics to the dry–wet (mediterranean) temperate zone via the arid interior, and encompassed substantial variation in all of the hypothesized controls of $N_{\text{area}}$ (Fig. 1). The AusPlots protocol involves sampling all species within a 100 × 100 m plot (White et al., 2012). We measured $N_{\text{area}}$, $^{13}$C, and LMA on all species at each site, and tested and quantified the effects of irradiance, $c_i : c_a$ ratio (from $^{13}$C), temperature, LMA, and N-fixation ability (26 % of the species sampled were N-fixers) on variation in $N_{\text{area}}$. The sampling design also allowed us to implement the trait gradient analysis method introduced by Ackerly and Cornwell (2007), which has been surprisingly little used to date. A growing body of field mea-
measurements shows extensive leaf-trait variation within species and plant functional types (PFTs) (Kattge et al., 2011; Meng et al., 2015). Trait gradient analysis allows trait variation to be partitioned into a component due to variation within species and a component due to species replacement.

2 Materials and methods

Our analyses are based on 442 leaf measurements representing all species found in a 100 m × 100 m plot at each of 27 sites on a broad north–south transect across Australia (Fig. 1). We performed a regression analysis to test the relationships of $N_{\text{area}}$ with mean annual temperature (MAT), irradiance, plant trait leaf mass per area (LMA), $c_i : c_a$ ratio, and N-fixation capacity. We also fitted a statistical model in which $N_{\text{area}}$ was treated as the sum of a metabolic component proportional to predicted (optimal) photosynthetic capacity at standard temperature (based on temperature, irradiance, and $c_i : c_a$ ratio) and a structural component proportional to LMA. Finally, we carried out a trait gradient analysis in order to quantify the contributions of environment vs. species identity to variation in $N_{\text{area}}$, $c_i : c_a$ ratio, and LMA.

2.1 Climate data and analysis

Climatological data for the 27 sites were obtained from the eMAST/ANUClimate data set (www.emast.org.au), which extends from 1970 to 2012 with 1 km spatial resolution across the entire continent. Mean annual precipitation (MAP) over this period at the sampling sites ranged from 154 to 1726 mm and mean annual temperature (MAT) from 14.1 to 27.6°C. The moisture index ($MI = P / E_{q}$, where $P$ is mean annual precipitation and $E_{q}$ is equilibrium evapotranspiration, calculated with the STASH program: Gallego-Sala et al., 2012) varied from 0.07 to 0.82. The mean incident flux of photosynthetically active radiation (PAR) during daylight hours, expressed as photosynthetic photon flux density ($\mu$mol m$^{-2}$ s$^{-1}$), was also calculated using STASH. This incident flux (at the top of the canopy) was averaged through the canopy using Beer’s law, as follows. First leaf area index ($L$) was estimated from a remotely sensed (MODIS NBAR- derived using MOD43A4: http://remote-sensing.nci.org.au/u39/public/html/modis/fractionalcover-clw) fractional cover of photosynthetic vegetation ($f_v$) at 1 km resolution at each site, from data assembled by the TERN AusCover facility (Guerschman et al., 2009):

$$L \approx -(1/k) \ln (1 - f_v),$$

(1)
where \( k = 0.5 \). Then absorbed PAR per unit leaf area (\( I_L \)) was calculated as
\[
I_L \approx I_0/(1 - e^{-kL})/L \approx I_0k f_v/\ln \left[1/(1 - f_v)\right],
\]
where \( I_0 \) is the incident PAR above the canopy. This calculation yields \( I_L \approx I_0 \) for sparse vegetation (\( L < 1 \)), but \( I_L \) becomes progressively smaller than \( I_0 \) as foliage density increases, reflecting the fact that the irradiance experienced by the average species is much lower in, say, a closed woodland than in an open shrubland, even if the PAR incident at the top of canopy is the same. In dense vegetation \( I_L \) will underestimate the PAR exposure of canopy dominants and overestimate the PAR exposure of understory species. However, the use of a canopy average in this way was a necessary approximation (because we did not have quantitative information about the canopy position of each species) and considered preferable to using \( I_0 \), which will systematically overestimate PAR exposure for most species in a dense community.

### 2.2 Foliage sampling and analysis

Mature outer-canopy leaves of each species were sampled during the growing season using the AusPlots methodology (White et al., 2012). (Note that in denser vegetation many species sampled are in the understory, so their “outer-canopy” leaves are still shaded by the overstory. Many species thus receive considerably reduced sunlight compared to the overstory, implying that the canopy-average irradiance \( I_L \) is more suitable than the top-of-canopy value \( I_0 \) as a community measure of irradiance.) In total, the 27 selected sites included 442 unique species, of which 37 were \( C_4 \) plants (not analysed further here). LMA was measured on the archived leaf samples by scanning and weighing the leaves. Subsamples (a mixture of material from at least two replicates) were analysed for C and N contents and bulk \( \delta^{13}C \) at the Stable Isotope Core Laboratory of Washington State University, USA. \( N_{\text{area}} \) was calculated from N content and LMA. Carbon isotope discrimination (\( \Delta \)) values were derived from the reported \( \delta^{13}C \) values using the standard formula
\[
\Delta = (\delta_{\text{air}} - \delta_{\text{plant}})/(1 + \delta_{\text{plant}}),
\]
where \( \delta_{\text{air}} \) is the carbon isotope composition of air and \( \delta_{\text{plant}} \) is the carbon isotope composition of the plant material. Because of the different diffusion rates and biochemical rates of carboxylation between \( ^{13}CO_2 \) and \( ^{12}CO_2 \), \( \Delta \) can be used to estimate the \( c_i : c_a \) ratio as
\[
c_i : c_a \approx (a + \Delta)/(b - a),
\]
where the recommended standard values are \( a = 4.4\%e \) and \( b = 27\%e \) (e.g. Cernusak et al. 2013).

### 2.3 Analysis of \( V_{\text{cmax}} \)

Values of \( V_{\text{cmax}} \) were predicted based on the co-ordination hypothesis, by equating the carboxylation- and electron transport-limited rates of photosynthesis and, as a simplifying assumption, treating the electron transport-limited rate as proportional to absorbed PAR (i.e. ignoring the saturation of the electron transport rate at high irradiances). These assumptions lead to the following estimate:
\[
V_{\text{cmax}} \approx \varphi_0 I_L (c_i + K)/(c_i + 2\Gamma^*),
\]
where \( \varphi_0 \) is the intrinsic quantum efficiency of photosynthesis (0.093; Long et al., 1993), \( c_i \) is the leaf-internal concentration of CO\(_2\), \( K \) is the effective Michaelis–Menten coefficient of Rubisco, and \( \Gamma^* \) is the photorespiratory compensation point. Values of both these quantities and their activation energies (governing their temperature responses) are based on the empirical in vivo determinations by Bernacchi et al. (2001) widely used in photosynthesis research. Both \( K \) and \( \Gamma^* \) were evaluated at standard atmospheric pressure and oxygen concentration, and site \( \text{MAT} \). Predicted values of \( V_{\text{cmax}} \) were adjusted to 25\(^\circ\)C, because the amount of N allocated to Rubisco and other enzymes involved in carboxylation should be proportional to \( V_{\text{cmax}} \) at a standard temperature, not at the growth temperature.

### 2.4 Statistical methods

All statistics were performed in R3.1.3 (R Core Team, 2015). Linear regressions were fitted using the \textit{lm} function, partial residual plots were generated using the \textit{visreg} package, and the relative contributions of different predictors were quantified using the \textit{relaimpo} package. In a first, exploratory statistical analysis, a linear model was fitted for \( \ln N_{\text{area}} \) with \( c_i : c_a \), MAT, \( \ln I_L \), \( \ln LMA \), and the factor “N-fixer” as predictors. The regression slopes of \( \ln N_{\text{area}} \) against \( c_i : c_a \), MAT and \( \ln I_L \) can all be independently predicted from the co-ordination hypothesis by differentiation of Eq. (5) (see Appendix A; note that these formulae explicitly predict the slopes for \( \ln N_{\text{area}} \)). These predicted values were compared with the fitted values and their 95\% confidence limits in order to assess support for the co-ordination hypothesis.

In a second analysis, community-mean values were calculated as simple averages across the species in each plot, omitting the factor “N-fixer”. A linear model was fitted to the community means of \( \ln N_{\text{area}} \) as a function of \( c_i : c_a \), MAT, \( \ln I_L \), and \( \ln LMA \) to assess the predictability of leaf N at the community level.

In a third analysis, \( N_{\text{area}} \) was modelled as a linear combination of the predictors Rubisco N, \( N_{\text{rubisco}} \) (derived from predicted \( V_{\text{cmax}} \) at 25\(^\circ\)C), and structural N, \( N_{\text{structure}} \) (derived from LMA using the empirical relationship \( N_{\text{structure}} = 10^{-2.67} \) LMA\(^{0.99} \), in g m\(^{-2} \); Yusuke Onoda, personal communication 2015), including “N-fixer” as a factor and allowing interactions of the predictors with this factor.
2.5 Trait gradient analysis

Trait gradients were generated for ln LMA, ln N_area, and c_i : c_a following the analysis method of Ackerly and Cornell (2007), again using simple averages across species to estimate community means. In this analysis species trait values were plotted against site-mean trait values. By definition, the regression of the species trait values against site-mean trait values has a slope of unity. For a perfectly plastic trait, regression of trait variation within species against the site-mean trait values would also yield a slope of unity. The common within-species slope that this approach provides is a measure of the fraction of trait variation due to phenotypic plasticity and/or genotypic variability. Its one-complement measures the fraction due to species turnover. Natural log transformation was applied to LMA and N_area because of their large variance and skewed distributions, but not to c_i : c_a, because of its small variance and approximately normal distribution.

3 Results

3.1 Leaf N variations with climate and leaf traits

Significant partial relationships were found for ln N_area vs. c_i : c_a, MAT, and ln I_L (Table 1, Fig. 2). The relationship was negative for c_i : c_a, as expected, because lower c_i : c_a implies that a greater photosynthetic capacity is required to achieve a given assimilation rate (or equivalently: a stronger CO_2 drawdown is enabled by a higher V_{cmax}). The relationship was also negative for MAT, as expected, because there is an inverse relationship between temperature and the quantity of leaf proteins required to support a given value of V_{cmax}. The relationship was positive for ln I_L (PAR), as expected, because the higher the irradiance, the greater the carboxylation capacity required for co-limitation with the rate of electron transport.

Theoretical slopes for these relationships (derived in Appendix A) are compared with the fitted slopes in Table 1. For ln N_area vs. ln I_L, the theoretical slope is unity. The fitted slope of 0.874 (95 % confidence limits: 0.685, 1.063) was statistically indistinguishable from unity. (A slope significantly greater than unity was found for ln N_area vs. ln I_0, i.e. top-of-canopy PAR, as expected, as this measure underestimates the change in mean canopy PAR along the gradient from sparse, high-PAR to dense, lower-PAR communities.)

For ln N_area vs. c_i : c_a, the fitted slope of −0.611 (−1.107, −0.115) was fortuitously close to the theoretical slope of −0.615, although the value was only weakly constrained for these data. For ln N_area vs. MAT, the theoretical slope was obtained by subtracting the “kinetic” slope of ln V_{cmax} vs. temperature (from the activation energy of carboxylation as given by Bernacchi et al., 2001) from the shallow positive slope implied by Eq. (5). The kinetic effect was dominant and results in an overall predicted negative slope of −0.048. The fitted slope of −0.047 (−0.060, −0.034) was indistinguishable from this theoretical slope, indicating acclimation to temperature by diminished allocation of N to metabolic functions at higher temperature, offsetting the increased reaction rate predicted by the Arrhenius equation. However, this slope was shallower than would be predicted by the Arrhenius equation alone, reflecting the reduced quantum efficiency of assimilation (a higher V_{cmax} is required to support a given assimilation rate) at higher temperatures.

The proportion of leaf N allocated to Rubisco has generally been found to decline, while the total N allocated to cell walls increases with increasing LMA (Hikosaka and Shigeno, 2009). Figure 2 shows a strong positive partial relationship between ln N_area and LMA. N-fixers had generally higher N_area than non-N-fixers (Fig. 2e: p<0.001). The predictors together explained 55 % of the variation in leaf N across species and sites.

Fully 82 % of the variation in the community-mean value of ln N_area could be explained by the combination of community-mean LMA and environmental variables. Significant partial relationships of community-mean ln N_area with MAT, ln I_L, and ln LMA (Table 2) were consistent with the results obtained at species level. The fitted slopes of ln N_area against ln I_L and MAT were again indistinguishable from the theoretical values, albeit with wide error bounds due to the much smaller sample size (27 as opposed to 405). The community-level partial relationship between ln N_area and c_i : c_a showed a negative slope as predicted, although this relationship was barely significant (p ≈ 0.1) due to the small sample size.

3.2 Leaf N as the sum of metabolic and structural components

Highly significant (p < 0.001) positive relationships were found between N_area and the predicted Rubisco-N content per unit leaf area (N_{rubisco}), and the predicted cell-wall N content per unit leaf area (N_{structure}) (Fig. 3). A priori we would expect the regression coefficient for N_{structure} to be close to unity, and that for N_{rubisco} to be about 6 to 20 (if Rubisco constitutes about 5 to 15 % of total leaf protein: Evans, 1989; Evans and Seemann, 1989; Onoda et al., 2004). The fitted slopes of 1.2 (p < 0.001; 95 % confidence limits: 1.0, 1.4) and 9.5 (p < 0.001; 7.6, 11.5) in Table 3 respectively were consistent with these expectations.

There was no significant main effect of the factor “N-fixer” and no significant interaction between N_{rubisco} and the factor “N-fixer”. The co-ordination hypothesis predicts that the metabolic component of N_area should be environmentally optimized and therefore independent of N supply. This could not be tested without direct measurements of V_{cmax} or N_{rubisco}, which were precluded by the design of this study. However, N-fixers showed a steeper relationship between N_area and N_{structure}. This was manifested as a significant inter-
action between the factor “N-fixing” and \( N_{\text{structure}} \) \((p < 0.01)\). This model, in which \( N_{\text{area}} \) was decomposed into a metabolic component predicted by the co-ordination hypothesis and a structural component proportional to LMA, explained 52% of the variance in \( N_{\text{area}} \) across species and sites. The relative importance of variations in the metabolic and structural components was determined to be 39 and 61% respectively, showing inter alia the importance of variation in LMA in determining leaf N content.

3.3 Quantifying trait plasticity vs. species turnover

In total, 243 C3 species were sampled at two or more sites. These species allowed calculation of a common slope, being an estimate of trait plasticity \textit{sensu lato} (that is, phenotypic plasticity or genetic adaptation or both) across species (Fig. 4), for the traits \( c_i : c_a \), \( \ln \text{LMA} \), and \( \ln \text{N}_{\text{area}} \). Contrasting results were obtained for the three traits. It appeared that \( c_i : c_a \) is perfectly plastic, with a common (within-species) slope indistinguishable from unity. The common slope of \( N_{\text{area}} \) was close to 0.5, indicating approximately equal contributions of plasticity and species turnover to the total variation. In the case of LMA, however, there was significant heterogeneity \((p < 0.05)\) among the within-species slopes, with \textit{Marsdenia viridiflora} showing a significantly steeper slope than the other species. After excluding this species, the common slope for LMA was also close to 0.5. A positive common slope indicates the ability of species to adapt their leaf morphology to environment. The positive common slope found for \( N_{\text{area}} \) is consistent with this trait’s nature as a combination of metabolic and structural components; its similar-

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Table 1. Linear regression coefficients for ln \( N_{\text{area}} \) (g m\(^{-2}\)) as a function of \( c_i : c_a \) (from \( \delta^{13}\text{C} \)), ln (mean canopy PAR, \( I_L \)) (\( \mu \text{mol m}^{-2} \text{s}^{-1} \)), MAT (\(^{\circ}\)), ln LMA (g m\(^{-2}\)), and the factor “N-fixing” at species level.

<table>
<thead>
<tr>
<th>Estimated</th>
<th>Predicted</th>
<th>( p )</th>
<th>Relative importance</th>
</tr>
</thead>
<tbody>
<tr>
<td>( c_i : c_a )</td>
<td>-0.611 ± 0.252</td>
<td>-0.615</td>
<td>&lt; 0.01</td>
</tr>
<tr>
<td>ln ( I_L )</td>
<td>0.874 ± 0.096</td>
<td>1</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>MAT</td>
<td>-0.047 ± 0.007</td>
<td>-0.048</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>ln LMA</td>
<td>0.415 ± 0.036</td>
<td>n/a</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>“N-fixing”</td>
<td>0.306 ± 0.041</td>
<td>n/a</td>
<td>&lt; 0.001</td>
</tr>
</tbody>
</table>

n/a: not applicable.
Figure 3. Partial residual plots for the linear regression of \(N_{\text{area}}\) as a function of independently predicted values of \(N_{\text{rubisco}}\) and \(N_{\text{structure}}\) (all in g m\(^{-2}\)) at species level. Blue: N-fixers; red: non-N-fixers.

Table 2. Linear regression coefficients for community-mean (simple average) values of ln \(N_{\text{area}}\) (g m\(^{-2}\)) as a function of \(c_l : c_a\) (from \(\delta^{13}\)C), ln (mean canopy PAR, \(I_L\)) (\(\mu\)mol m\(^{-2}\) s\(^{-1}\)), MAT (°), and ln LMA (g m\(^{-2}\)).

<table>
<thead>
<tr>
<th>Estimated</th>
<th>Predicted</th>
<th>(p)</th>
<th>Relative importance</th>
</tr>
</thead>
<tbody>
<tr>
<td>(c_l : c_a)</td>
<td>-1.60 ± 0.94</td>
<td>-0.615</td>
<td>n.s.</td>
</tr>
<tr>
<td>ln (I_L)</td>
<td>0.70 ± 0.23</td>
<td>1</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>MAT</td>
<td>-0.035 ± 0.016</td>
<td>-0.048</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>ln LMA</td>
<td>0.57 ± 0.19</td>
<td>n/a</td>
<td>&lt;0.001</td>
</tr>
</tbody>
</table>

n.s.: no significance. n/a: not applicable.

The variety of environments provided in this study by the long transect, and the number of species sampled, allowed us to statistically separate the effects of \(c_l : c_a\), irradiance, temperature, and LMA on \(N_{\text{area}}\). The relationships with \(c_l : c_a\), irradiance, and temperature were in the directions and magnitudes predicted by the co-ordination hypothesis. The relationship with site mean irradiance had a slope as predicted by the co-ordination hypothesis (i.e. close to 1), but a strong relationship, with a steeper slope as expected, was found when top-of-canopy irradiance was used instead of the canopy mean – indicating that both spatial variations and within-canopy shading were contributing to the relationship with site mean irradiance. We performed an additional regression using leaf nitrogen content per unit mass (\(N_{\text{mass}}\)) which showed, as expected, identical fitted coefficients for all predictors except LMA (Appendix B). However, because of the regression coefficient of ln \(N_{\text{area}}\) with respect to ln LMA < 1, the regression coefficient of ln \(N_{\text{mass}}\) with respect to ln LMA < 0, i.e. \(N_{\text{mass}}\), declines with increasing LMA – as has been widely reported. We also tried a regression of \(N_{\text{mass}}\) on the same set of predictors but without the inclusion of LMA; this yielded a much poorer fit and is not shown.

High \(N_{\text{area}}\) in plants from arid environments has been described often, and has traditionally been explained as a consequence of high N supply in environments with low rainfall (reducing leaching losses) and restricted plant cover (reducing total vegetation N demand) (e.g. Field and Mooney, 1986). This explanation would imply that plants in wetter environments have lower (and suboptimal) \(N_{\text{area}}\) due to low availability of N. However, the negative relationship commonly found between \(c_l : c_a\) and \(N_{\text{area}}\) supports an alternative, adaptive (plant-centred) explanation. The least-cost hypothesis (Wright et al., 2003; Prentice et al., 2014) predicts lower \(c_l : c_a\) in drier environments. This is because the drier the atmosphere, the greater the flux of water required to support a given rate of assimilation, which in turn shifts the balance of costs and benefits towards investment in photosynthetic capacity (\(V_{c_{\text{max}}}\)) and away from water transport capacity. When \(c_l : c_a\) is lower, the co-ordination hypothesis predicts that a higher \(V_{c_{\text{max}}}\) (and therefore higher \(N_{\text{area}}\)) will be optimal, in order for the leaves to fully utilize the available light. The co-ordination hypothesis also predicts a further in-
Figure 4. Trait means and regression lines for all 243 C3 plant species in the 27 study sites. Note the logarithmic scales for \( N_{area} \) (g m\(^{-2}\)) and LMA (g m\(^{-2}\)). Thin red dashed lines represent individual within-species regression lines of non-N-fixer species. Thin blue lines represent individual within-species regression lines of N-fixer species. The black dashed line represents the overall regression line, which has a slope of unity by definition. Grey dots denote individual species–site combinations. Common within-species slopes are 0.53 ± 0.11 (ln \( N_{area} \)), 1.02 ± 0.12 (\( c_1 : c_a \)), and 0.55 ± 0.11 (ln LMA).

Table 3. Linear regression coefficients for \( N_{area} \) as a function of independently predicted values of \( N_{rubisco} \) and \( N_{structure} \) (all in g m\(^{-2}\)) at species level.

<table>
<thead>
<tr>
<th></th>
<th>Estimated</th>
<th>Predicted</th>
<th>( p )</th>
<th>Relative importance</th>
<th>( R^2 )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( N_{rubisco} )</td>
<td>9.5 ± 2.0</td>
<td>6–20</td>
<td>&lt;0.001</td>
<td>39 %</td>
<td></td>
</tr>
<tr>
<td>( N_{structure} )</td>
<td>1.2 ± 0.2</td>
<td>1</td>
<td>&lt;0.001</td>
<td>61 %</td>
<td>52 %</td>
</tr>
<tr>
<td>( N_{structure}; \text{“N-fixer”} )</td>
<td>1.0 ± 0.3</td>
<td>n/a</td>
<td>&lt;0.01</td>
<td>n/a</td>
<td></td>
</tr>
</tbody>
</table>

n/a: not applicable.

Increase in \( N_{area} \) with increasing aridity due to reduced cloudiness and reduced shading by competitors, both factors tending to increase \( I_L \) (and both apparently contributing to the fitted relationship of \( N_{area} \) with \( I_L \)). Thus the co-ordination hypothesis could account for independent positive effects of site irradiance and aridity on \( N_{area} \), as previously reported by Wright et al. (2005). The fitted relationship of \( N_{area} \) with temperature, PAR, and \( c_1 : c_a \) is consistent with our theoretical prediction, which implicitly includes all of these effects.

Despite the large within-site variation in LMA found at all points along the aridity gradient, there is a significant tendency for LMA to increase with aridity, perhaps because of the resistance to dehydration conferred by stiffer leaves (Niinemets, 2001; Wright and Westoby, 2002; Harrison et al., 2010) and/or the need for leaves to avoid overheating under transient conditions of high radiation load and low transpiration rates combined with low wind speed (Leigh et al., 2012). This increase in LMA is inevitably accompanied by an increasing structural N component.

Thus, several distinct aspects of plant allocation tend to increase \( N_{area} \) along gradients of increasing dryness. The predicted response of \( N_{rubisco} \) to temperature is a result of opposing effects: the declining efficiency of photosynthesis with increasing temperature (due to the temperature dependencies of \( K \) and \( \Gamma^* \)) is offset by the increased catalytic capacity of Rubisco at higher temperatures. The latter effect is predicted to be stronger, implying reduced \( N_{area} \) with increasing temperature, as observed.

4.2 The predictability of leaf N

Predicted \( N_{rubisco} \) and \( N_{structure} \) together explained more than half of the variation in total \( N_{area} \) across species and sites. Our approach to predicting these two quantities invokes a simplified formula, Eq. (5), which is based on the co-ordination hypothesis for \( N_{rubisco} \), assuming proportionality with Rubisco capacity, and assumes a simple proportionality with LMA for \( N_{structure} \). Our finding of highly significant multiple regression coefficients for both variables indicates that the prediction obtained when taking both into account is more accurate than could be obtained from either variable alone. Osnas et al. (2013), analysing a large global leaf-trait data set and applying a novel method to determine the extent to which different traits are area- vs. mass-
proportional, found leaf N to be an intermediate case. This is to be expected if leaf N is, as our results suggest, a composite of an area-proportional (Nrubisco) and a mass-proportional (Nstructure) component. The two predictors (Rubisco capacity and LMA) are not fully independent, because leaves with higher photosynthetic capacity tend to have higher LMA for structural reasons. But such leaves must have increased structural N as well. By showing independently significant regression coefficients for modelled Nrubisco and LMA, the multiple regression results establish that successful prediction of Narea requires consideration of both components, and that each has an independent effect, irrespective of their correlation ($r^2 = 0.28$ in this data set). Osnas et al. (2013) also fitted various statistical models for the relationships among leaf traits. Their “model LN” for ln Narea vs. ln LMA yielded a slope of 0.38 (95% confidence interval 0.36 to 0.40). This value, based on a global data set, can be compared directly with – and is indistinguishable from – our fitted partial regression coefficient of ln Narea vs. ln LMA, which is 0.42 (0.34 to 0.49) (Table 1).

In reality, however, leaf N does not consist exclusively of Rubisco and cell-wall constituents. Leaf N includes multiple additional components, including other photosynthetic proteins, proteins of the light-harvesting complexes and electron transport chains, cytosolic proteins, ribosomes and mitochondria, nucleic acids (which account for about 10–15% of leaf N; Chapin III and Kedrowski, 1983), and N-based defensive compounds. It is possible that the higher N found for N-fixers resides in N-based osmolytes (Erskine et al., 1996) or defence compounds (Gutschick, 1981). Nonetheless, our simplifications suggest that Narea – especially at the community level, which is key for large-scale modelling – is, to first order, inherently predictable from leaf morphology and the physical environment. A corollary is that limitation in N supply may act primarily by changing plant allocation patterns (reducing allocation to light capture by leaves while increasing allocation to N uptake by roots), rather than by altering leaf stoichiometry.

4.3 Trait variations within and between species

By testing for acclimation along spatial gradients, the design of our study did not allow phenotypic plasticity to be distinguished from genetic adaptation. Phenotypic plasticity is the ability of a genotype to alter its expressed trait values in response to environmental conditions (Bradshaw, 1965; Sultan, 2000). A part of the observed variation in trait values within species could be due to shifts in the occurrence and frequency of different genotypes, producing different preferred trait values. Thus, when we refer to traits as “plastic”, this should be understood in a broad sense to allow the possibility of a genetic component of the observed adaptive differentiation within species. Seasonal acclimation within individual plants can provide more direct evidence for phenotypic plasticity (Togashi et al., 2017), whereas in this study we disregard possible seasonal variations and instead relate trait variations to the mean annual environment. However, by sampling all of the species present at each site and including measurements on species at multiple sites, we could distinguish between the contribution of plasticity sensu lato (phenotypic plasticity and/or genetic adaptation) vs. species turnover, i.e. the progressive replacement of species with different mean trait values, to spatial variation in the community-mean values of a given trait. We found that $\delta^{13}$C was perfectly plastic, perhaps not surprisingly, as variations in $c_i : c_a$ are under stomatal control. In contrast, LMA and Narea showed approximately equal contributions from plasticity and species turnover.

4.4 Implications for modelling

There has been a surge of interest in schemes to predict continuous trait variation in DGVMs (e.g. Scheiter et al., 2013; Fyllas et al., 2014; van Bodegom et al., 2014; Ali et al., 2015; Fisher et al., 2015; Meng et al., 2015; Sakschewski et al., 2015). Some trait-based modelling approaches have relied on empirical information on trait–trait and trait–environment covariation, but others (e.g. Scheiter et al., 2013) have aimed to represent the adaptive nature of trait variation explicitly. Our focus has been on testing an explicit adaptive hypothesis for the controls of one key trait, Narea, which in addition to a structural component (necessarily linked to LMA) includes an important metabolic component, reflecting the leaf-level investment in photosynthetic proteins. All models that attempt to represent the coupling between C and N cycles in terrestrial ecosystems require a method to calculate leaf N content, given other environmental and plant characteristics. Some models prescribe fixed values of $V_{\text{cmax}}$ (per plant functional type), but this approach does not take account of the observed variation in $V_{\text{cmax}}$ with environmental conditions. Models that assume proportionality between $V_{\text{cmax}}$ and Narea neglect the important variation in leaf structural N. We have shown that Narea is predictable, to a degree that is useful for modelling, when both metabolic and structural components are taken into account. Our prediction is based on LMA, $c_i : c_a$, and a theoretically predicted value of $V_{\text{cmax}}$ based on the co-ordination hypothesis – for which there is strong independent evidence (e.g. Maire et al., 2012). The partial responses of Narea to $c_i : c_a$, irradiance, and temperature are consistent with predictions of the co-ordination hypothesis, and the inclusion of predicted $V_{\text{cmax}}$ adds significantly and substantially to the predictive power of LMA and $c_i : c_a$ alone. As both LMA (Wright et al., 2005) and $c_i : c_a$ (Prentice et al., 2014) show relationships with environment, our results suggest a possible route towards a general adaptive scheme for the prediction of major leaf traits in DGVMs, which would be an improvement on models that assume a one-to-one relationship between photosynthetic capacity and Narea (see e.g. Adams et al., 2016, who showed that there is considerable variation in Narea among N-fixers that is unrelated to photosynthetic capacity). Our results also suggest
some priorities for trait data collection and analysis: to test the predicted controls of $N_{\text{area}}$ over a wider range of environments, and to test the predicted environmental controls of $V_{c_{\text{max}}}$ directly in the field.

Our application of trait gradient analysis also points out a way towards process-based treatments of functional trait diversity in next-generation models. It is increasingly accepted that models could, and should, sample “species” from continuous gradients of traits rather than fix the traits associated with discrete PFTs. A hybrid approach to modelling $N_{\text{area}}$ based on the present analysis would consider $N_{\text{area}}$ explicitly as the sum of metabolic and structural components. The metabolic component would be treated as plastic and subject to environmental optimization (in space and time), consistent with the least-cost and co-ordination hypotheses. The structural component would be tied to LMA, which is a key variable of the “leaf economics spectrum” (Wright et al., 2004), strongly expressed both within and between environments and therefore requiring a broad range of values to be assigned to model “species”.

Finally, we note that if our results can be corroborated more widely, this would point to the need for a shift in the way N “limitation” is treated – both in models and in analyses of field data. In studies of the relationship between $V_{c_{\text{max}}}$ and leaf N, for example, it is conventional to plot N on the $x$ axis and $V_{c_{\text{max}}}$ on the $y$ axis, and it is then often stated that the positive relationship found shows that variation in leaf N “causes” variation in $V_{c_{\text{max}}}$. But all that is shown on the graph is a correlation, and our “plant-centred” interpretation is the opposite of the conventional one: that is, $V_{c_{\text{max}}}$ is adaptively matched (acclimated) to environmental conditions, and the metabolic component of leaf N is a consequence of this acclimation. Low N availability would then result in reduced allocation of C (and N) to leaves, and increased allocation below ground – which is also an adaptive response, but at the whole-plant rather than the leaf level.
Appendix A: Theoretical responses of N_{area} to environmental predictors

We estimate optimal V_{cmax} by \( \gamma_0 I_L(c_i + K)/(c_i + 2\Gamma^*) \) (Eq. 5). Holding other variables constant, the sensitivity of this estimate to absorbed PAR is given by the derivative of its natural logarithm with respect to \( \ln I_L \):

\[
\frac{\partial \ln V_{cmax}}{\partial \ln I_L} = 1. \tag{A1}
\]

Similarly, the sensitivity of this estimate to \( c_i \) is given by

\[
\frac{\partial \ln V_{cmax}}{\partial c_i} = \frac{(2\Gamma^* - K)}{[(c_i + K)(c_i + 2\Gamma^*)]} \tag{A2}
\]

and its sensitivity to the \( c_i : c_a \) ratio is smaller than this by a factor \( c_a \).

Temperature-dependent reaction rates are described by the Arrhenius equation:

\[
\ln x(T) - \ln x(T_{ref}) = \frac{\Delta H}{R} \left( \frac{1}{T_{ref}} - \frac{1}{T} \right), \tag{A3}
\]

where \( x \) is the rate parameter of interest, \( T \) is the measurement temperature (K), \( T_{ref} \) is the reference temperature (here 298 K), \( \Delta H \) is the activation energy of the reaction (J mol\(^{-1}\) K\(^{-1}\)), and \( R \) is the universal gas constant (8.314 J mol\(^{-1}\) K\(^{-1}\)). Linearizing Eq. (A3) around \( T_{ref} \) yields

\[
\ln x(T) - \ln x(T_{ref}) \approx \frac{\Delta H}{R T_{ref}^2} \Delta T, \tag{A4}
\]

where \( \Delta T = T - T_{ref} \). Thus, from Eq. (5),

\[
\ln V_{cmax}^{25} \approx \ln V_{cmax} - \frac{\Delta H_v}{R T_{ref}^2} \Delta T, \tag{A5}
\]

where \( \Delta H_v \) is the activation energy of \( V_{cmax} \). The sensitivity of \( V_{cmax}^{25} \) to \( T \) is then

\[
\frac{\partial \ln V_{cmax}^{25}}{\partial T} = \frac{\partial \ln V_{cmax}}{\partial T} - \frac{\Delta H_v}{R T_{ref}^2}
\]

\[
\frac{\partial \ln V_{cmax}^{25}}{\partial T} = \frac{\partial K}{\partial T}(c_i + K) - 2(\partial \Gamma^*/\partial T)/
\]

\[
(c_i + 2\Gamma^*) - \frac{\Delta H_v R}{T_{ref}^2}, \tag{A6}
\]

where \( K = K_c (1 + O/K_o) \), hence

\[
\frac{\partial K}{\partial T} = \frac{\partial K_c}{\partial T}
\]

\[
+ [(\partial K_c/\partial T) K_o - (\partial K_o/\partial T) K_c] O/K_o^2, \tag{A7}
\]

where \( O \) is the atmospheric concentration of oxygen and \( \Gamma^* \) and the Michaelis–Menten coefficients for carboxylation (\( K_c \)) and oxygenation (\( K_o \)) respectively have values at \( T_{ref} \) (in \( \mu \)mol mol\(^{-1}\)) and activation energies as given by Bernacchi et al. (2001).
Appendix B: Partial responses of $N_{\text{mass}}$ to environmental predictors

![Graphs](image.png)

Figure B1. Partial residual plots for the regression of ln ($N_{\text{mass}} \times 100$) (g g$^{-1}$) as a function of $c_i : c_o$ (from $\delta^{13}$C), ln (mean canopy PAR, IL) ($\mu$mol m$^{-2}$ s$^{-1}$), MAT ($^\circ$C), ln LMA (g m$^{-2}$), and the factor “N-fixer” at species level.

Table B1. Linear regression coefficients for ln ($N_{\text{mass}} \times 100$) (g g$^{-1}$) as a function of $c_i : c_o$ (from $\delta^{13}$C), ln (mean canopy PAR, $I_L$) ($\mu$mol m$^{-2}$ s$^{-1}$), MAT ($^\circ$C), ln LMA (g m$^{-2}$), and the factor “N-fixer” at species level. Note that $N_{\text{mass}}$ was multiplied by 100 before logarithmic transformation.

<table>
<thead>
<tr>
<th></th>
<th>Estimated</th>
<th>Predicted</th>
<th>$p$</th>
<th>$R^2$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$c_i : c_o$</td>
<td>$-0.611 \pm 0.252$</td>
<td>$-0.615$</td>
<td>$&lt;0.01$</td>
<td></td>
</tr>
<tr>
<td>ln $I_L$</td>
<td>$0.874 \pm 0.096$</td>
<td>1</td>
<td>$&lt;0.001$</td>
<td></td>
</tr>
<tr>
<td>MAT</td>
<td>$-0.047 \pm 0.007$</td>
<td>$-0.048$</td>
<td>$&lt;0.001$</td>
<td></td>
</tr>
<tr>
<td>ln LMA</td>
<td>$-0.585 \pm 0.036$</td>
<td>n/a</td>
<td>$&lt;0.001$</td>
<td>51%</td>
</tr>
<tr>
<td>“N-fixer”</td>
<td>$0.306 \pm 0.041$</td>
<td>n/a</td>
<td>$&lt;0.001$</td>
<td></td>
</tr>
</tbody>
</table>

n/a: not applicable.
Information about the Supplement
Species analyzed in this study can be found in Supplement S1.

The Supplement related to this article is available online at doi:10.5194/bg-14-481-2017-supplement.

Author contributions. Iain Colin Prentice, Ning Dong, and Andrew J. Lowe planned and designed the study; Ning Dong carried out all the field measurements and performed the data analyses. Ning Dong and Iain Colin Prentice wrote the first draft; Bradley J. Evans supported the study through provision of climate data; Ian J. Wright assisted with data interpretation, contributed with ideas throughout, and suggested important improvements to the text. Stefan Caddy-Retalic contributed important ideas to improve the text. All authors contributed to subsequent versions.

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